No books or notes allowed. No laptop or wireless devices allowed. Show all your work for full credit. Write clearly and legibly.

Name (print): $\qquad$

| Question: | 1 | 2 | 3 | 4 | 5 | Total |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Points: | 40 | 20 | 30 | 20 | 0 | 110 |
| Score: |  |  |  |  |  |  |


| Question: | 1 | 2 | 3 | 4 | 5 | Total |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Bonus Points: | 0 | 0 | 0 | 0 | 20 | 20 |
| Score: |  |  |  |  |  |  |

Question 1....................................................................................... 40 point
Let $X$ be a continuous r.v. with p.d.f. given by

$$
f(x)= \begin{cases}p+2(1-p) x & 0<x<1 \\ 0 & \text { otherwise }\end{cases}
$$

where $p$ is a parameter.
(a) (10 points) for which value of $p$ is $f$ a valid p.d.f. (Hint: remember that there are 2 conditions you should check.)

Solution: We have

$$
\int_{0}^{1} f(x) d x=1
$$

for every $p$ so that we just need to check that $f(x)>0$. Since $f$ is linear in $x$ it is enough that $f(0)>0$ and $f(1)>0$. Thus we get

$$
0 \leq p \leq 2 .
$$

(b) (15 points) Compute the expected value $\mathbb{E}(X)$ and the variance $V(X)$ of $X$.

Solution: We have

$$
\mathbb{E}(X)=\int_{0}^{1} x f(x) d x=p \int_{0}^{1} x d x+2(1-p) \int_{0}^{1} x^{2} d x=\frac{p}{2}+\frac{2}{3}(1-p)=\frac{2}{3}-\frac{p}{6}
$$

and

$$
\mathbb{E}\left(X^{2}\right)=\int_{0}^{1} x^{2} f(x) d x=p \int_{0}^{1} x^{2} d x+2(1-p) \int_{0}^{1} x^{3} d x=\frac{p}{3}+\frac{1}{2}(1-p)=\frac{1}{2}-\frac{p}{6}
$$

so that

$$
V(x)=\frac{1}{36}(p(2-p)+2)
$$

(c) (15 points) Show that

$$
\mathbb{E}\left(\left(X-\frac{2}{3}\right)^{2}\right) \geq\left(\frac{p}{6}\right)^{2}
$$

(Hint: use Jensen inequality.)
Solution: Use Jensen inequality to get

$$
\mathbb{E}\left(\left(X-\frac{2}{3}\right)^{2}\right) \geq\left(\mathbb{E}(X)-\frac{2}{3}\right)^{2}=\left(\frac{1}{2}-\frac{p}{6}-\frac{1}{2}\right)^{2}
$$


Let $N_{k}, k=1,2,3, \ldots$, be an infinite sequence of geometric random variable with parameter $p_{k}=\frac{\lambda}{k}$, that is

$$
\mathbb{P}\left(N_{k}=n\right)=\left(1-p_{k}\right)^{n-1} p_{k} \quad \text { for } n \geq 1
$$

and $\mathbb{P}\left(N_{k}=n\right)=0$ for $n<1$. Moreover let $Y$ be an exponential r.v. with parameter $\lambda$, that is

$$
f_{Y}(y)=\lambda e^{-\lambda y} \quad \text { for } y \geq 0
$$

and $f_{Y}(y)=0$ for $y<0$.
Show that $Z_{k}=N_{k} / k$ converge in distribution to $Y$ as $k \rightarrow \infty$. (Hint: compute the c.d.f. of $Z_{k}$, that is $F_{k}(x)=\mathbb{P}\left(Z_{k} \leq x\right)$ for every real number $x$.)

Solution: Observe that we have

$$
\mathbb{P}\left(Z_{k} \leq x\right)=\mathbb{P}\left(N_{k} \leq k x\right)=\sum_{i=1}^{\lfloor k x\rfloor}\left(1-p_{k}\right)^{i-1} p_{k}=1-\left(1-p_{k}\right)^{\lfloor k x\rfloor-1}=\left(1-\frac{\lambda}{k}\right)^{\lfloor k x\rfloor-1}
$$

while

$$
\mathbb{P}(Y \leq y)=1-e^{-\lambda y}
$$

Observe that

$$
\frac{\lfloor k x\rfloor-1}{k} \rightarrow x
$$

as $k \rightarrow \infty$ so that

$$
\left(1-\frac{\lambda}{k}\right)^{\lfloor k x\rfloor-1}=\left(\left(1-\frac{\lambda}{k}\right)^{k}\right)^{\frac{\lfloor k x\rfloor-1}{k}} \rightarrow_{k \rightarrow \infty} e^{-\lambda x}
$$

for every $x$.

## Question 3

 30 pointA student is attempting a multiple choices exam. For each question there are 4 possible answers. He has a probability of 0.75 of knowing the correct answer. If he does not know the answer he chooses one answer uniformly and randomly. All questions and answers are independent.

To get a B he need to answer correctly $85 \%$ of the questions while to get an A he needs to answer correctly $95 \%$ of the questions.
(a) (15 points) If the test contains 40 questions, use a normal approximation (CLT) and the table provided to compute the probability $p_{B}$ that the student will get at least a B and the probability $p_{A}$ that the student will get a A .

Solution: Let $p$ be the probability that the student give a correct answer. We have

$$
p=0.75+0.25 \cdot 0.25=0.8125
$$

Let $X_{i}$ be 1 if he answer correctly to the $i$-th question and 0 otherwise. Thus $\mathbb{E}\left(X_{i}\right)=0.8125$ and $V\left(X_{i}\right)=0.1523$. We get

$$
\begin{aligned}
p_{B} & =\mathbb{P}\left(\sum_{i=1}^{40} X_{i}>0.85 \cdot 40\right)= \\
& =\mathbb{P}\left(\frac{\sum_{i=1}^{40} X_{i}-0.8125 \cdot 40}{0.390 \sqrt{40}}>\frac{(0.85-0.8125) \cdot 40}{0.390 \sqrt{40}}\right)= \\
& =1-\Phi(0.60)=0.274
\end{aligned}
$$

while

$$
\begin{aligned}
p_{B} & =\mathbb{P}\left(\sum_{i=1}^{40} X_{i}>0.95 \cdot 40\right)= \\
& =\mathbb{P}\left(\frac{\sum_{i=1}^{40} X_{i}-0.8125 \cdot 40}{0.390 \sqrt{40}}>\frac{(0.95-0.8125) \cdot 40}{0.390 \sqrt{40}}\right)= \\
& =1-\Phi(2.23)=0.013 .
\end{aligned}
$$

(b) (15 points) Let $p_{B}$ be the probability that a student that knows $75 \%$ of the answers will get a B or more. If the teacher wants $p_{B}$ to be less than 0.025 , how many question should there be on the exam.

Solution: He wants to find $N$ such that

$$
\mathbb{P}\left(\sum_{i=1}^{N} X_{i}>0.85 \cdot N\right) \leq 0.025
$$

This means

$$
\mathbb{P}\left(\frac{\sum_{i=1}^{N} X_{i}-0.8125 \cdot N}{0.390 \sqrt{N}}>\frac{(0.85-0.8125) \cdot \sqrt{N}}{0.390}\right)=1-\Phi(0.096 \cdot \sqrt{N}) \leq 0.025
$$

From the table we

$$
\Phi(1.96)=0.975
$$

so that he needs

$$
N>\left(\frac{1.96}{0.096}\right)^{2}=416
$$

questions.

Let $Z$ be a standard normal random variable. Find the $n$-th moment $m_{n}=\mathbb{E}\left(Z^{n}\right)$ of $Z$, for every $n$. (Hint: you can use integration by part to relate $m_{n}$ with $m_{n-2}$ and then use induction. Alternatively you can use the Taylor expansion around 0 of the m.g.f of a normal standard.)

Solution: By symmetry we just need to look at even $n$, that is $n=2 k$.
First method: integrating by part we get

$$
\begin{aligned}
m_{2 k} & =\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{\infty} z^{2 k} e^{-\frac{z^{2}}{2}} d z=\frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{\infty} z^{2 k-1} z e^{-\frac{z^{2}}{2}} d z= \\
& =(2 k-1) \frac{1}{\sqrt{2 \pi}} \int_{-\infty}^{\infty} z^{2 k-2} e^{-\frac{z^{2}}{2}} d z=(2 k-1) m_{2 k-2}
\end{aligned}
$$

we know that $m_{0}=1$ so that

$$
m_{2 k}=(2 k-1)(2 k-3)(2 k-5) \cdots 5 \cdot 3 \cdot 1
$$

Second method: We know that the moment generating function of $Z$ is

$$
M_{Z}(t)=e^{\frac{t^{2}}{2}}=\sum_{k=0}^{\infty} \frac{t^{2 k}}{2^{k} k!}
$$

so that

$$
\left.\frac{d^{k}}{d t^{k}} M_{Z}(t)\right|_{t=0}=\frac{(2 k)!}{2^{k} k!}=(2 k-1)(2 k-3)(2 k-5) \cdots 5 \cdot 3 \cdot 1
$$

5. (20 points (bonus)) Let $N_{1}, N_{2}$ and $N_{3}$ be discrete random variables with joint probability mass function

$$
p\left(n_{1}, n_{2}, n_{3}\right)=\mathbb{P}\left(N_{1}=n_{1} \& N_{2}=n_{2} \& N_{3}=n_{3}\right)=\frac{3^{-N} N!}{n_{1}!n_{2}!n_{3}!}
$$

if $n_{1}+n_{2}+n_{3}=N$ and 0 otherwise.
Compute the marginal mass function $p_{N_{1}}$ of $N_{1}$, that is

$$
p_{N_{1}}\left(n_{1}\right)=\mathbb{P}\left(N_{1}=n_{1}\right)
$$

and the conditional mass function $p_{N_{2}, N_{3} \mid N_{1}}$ of $N_{2}$ and $N_{3}$ given $N_{1}$, that is

$$
p_{N_{2}, N_{3} \mid N_{1}}\left(n_{2}, n_{3} \mid n_{1}\right)=\mathbb{P}\left(N_{2}=n_{2} \& N_{3}=n_{3} \mid N_{1}=n_{1}\right)
$$

(Hint: you can answer the question without doing any computation. Think what situation is described by $N_{1}, N_{2}$ and $N_{3}$.)

Solution: Observe that $N_{1}, N_{2}$ and $N_{3}$ are the result of repeating an experiment with 3 possible equiprobable outcomes (say $1,2,3) N$ times. $\mathbb{P}\left(N_{1}=n_{1}\right)$ represents the probability of obtaining $n_{1} 1 \mathrm{~s}$ when the probability of a 1 in $1 / 3$. Thus $\mathbb{P}\left(N_{1}=n_{1}\right)$ is a binomial with $p=1 / 3$ that is

$$
\mathbb{P}\left(N_{1}=n_{1}\right)=\frac{N!}{\left(N-n_{1}\right)!n_{1}!}\left(\frac{1}{3}\right)^{n_{1}}\left(\frac{2}{3}\right)^{N-n_{1}}
$$

On the other hand if you know you had exactly $n_{1} 1$ 's, the remaining outcomes are 2 or 3 , with equal probability. Thus

$$
p_{N_{2}, N_{3} \mid N_{1}}\left(n_{2}, n_{3} \mid n_{1}\right)=\frac{2^{-\left(N-n_{1}\right)}\left(N-n_{1}\right)!}{n_{2}!n_{3}!}
$$

if $n_{2}+n_{3}=N-n_{1}$ and 0 otherwise.
In formulas we have

$$
\begin{aligned}
p_{N_{1}}\left(n_{1}\right) & =\sum_{n_{2}, n_{3}} p\left(n_{1}, n_{2}, n_{3}\right)=\sum_{n_{2}+n_{3}=N-n_{1}} \frac{3^{-N} N!}{n_{1}!n_{2}!n_{3}!}= \\
& =\frac{3^{-N} 2^{N-n_{1}} N!}{\left(N-n_{1}\right)!n_{1}!} \sum_{n_{2}+n_{3}=N-n_{1}} \frac{2^{-\left(N-n_{1}\right)}\left(N-n_{1}\right)!}{n_{2}!n_{3}!}= \\
& =\frac{N!}{\left(N-n_{1}\right)!n_{1}!}\left(\frac{1}{3}\right)^{n_{1}}\left(\frac{2}{3}\right)^{N-n_{1}}
\end{aligned}
$$

so that $N_{1}$ is a binomial r.v. with $N$ trials and $p=1 / 3$.

Moreover we have

$$
\begin{aligned}
p_{N_{2}, N_{3} \mid N_{1}}\left(n_{2}, n_{3} \mid n_{1}\right) & =\frac{3^{-N} N!}{n_{1}!n_{2}!n_{3}!}\left(\frac{N!}{\left(N-n_{1}\right)!n_{1}!}\left(\frac{1}{3}\right)^{n_{1}}\left(\frac{2}{3}\right)^{N-n_{1}}\right)^{-1}= \\
& =\frac{2^{-\left(N-n_{1}\right)}\left(N-n_{1}\right)!}{n_{2}!n_{3}!}
\end{aligned}
$$

if $n_{2}+n_{3}=N-n_{1}$ and 0 otherwise.
Thus $N_{2}$ is a binomial r.v with $N-n_{1}$ trials and $p=1 / 2$.

## Useful Formulas

- Normal Distribution: if $Z$ is a standard normal r.v. then its density function is

$$
f(z)=\frac{1}{\sqrt{2 \pi}} e^{-\frac{z^{2}}{2}}
$$

while $E(Z)=0$ and $V(Z)=1$. The moment generating function $M_{Z}(t)$ is given by

$$
M_{Z}(t)=e^{\frac{t^{2}}{2}}
$$

Moreover

$$
\Phi(z)=\mathbb{P}(Z \leq z)
$$

is given in the table on next page. Finally if $X$ is normal with $\mathbb{E}(X)=\mu$ and $V(X)=\sigma^{2}$ then

$$
Y=\frac{X-\mu}{\sigma}
$$

is normal standard.

- Jensen's Inequality: If $X$ is a r.v. and $g$ is a convex function then

$$
\mathbb{E}(g(X)) \geq g(\mathbb{E}(X))
$$

- CLT: if $X_{i}$ is a sequence of i.i.d. random variable with expected value $\mu$ and variance $\sigma^{2}$ and

$$
S_{n}=\frac{1}{\sqrt{n}} \sum_{i=1}^{n} \frac{X_{i}-\mu}{\sigma}
$$

then $S_{n}$ converges in distribution to a normal standard r.v. $Z$.

- Convergence in Distribution: we say that the sequence $X_{n}$ converge in distribution to $X$ if

$$
\mathbb{P}\left(X_{n} \leq x\right) \rightarrow_{n \rightarrow \infty} \mathbb{P}(X \leq x)
$$

for every $x \in \mathbb{R}$.

## Table 1b: Standard Normal Probabilities

The values in the table below are cumulative probabilities for the standard normal distribution $Z$ (that is, the normal distribution with mean 0 and standard deviation 1). These probabilities are values of the following integral:

$$
P(Z \leq z)=\int_{-\infty}^{z} \frac{1}{\sqrt{2 \pi}} e^{-x^{2} / 2} d x
$$

Geometrically, the values represent the area to the left of $z$ under the density curve of the standard normal distribution:


| z | . 00 | . 01 | . 02 | . 03 | . 04 | . 05 | . 06 | . 07 | . 08 | . 09 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.0 | 0.5000 | 0.5040 | 0.5080 | 0.5120 | 0.5160 | 0.5199 | 0.5239 | 0.5279 | 0.5319 | 0.5359 |
| 0.1 | 0.5398 | 0.5438 | 0.5478 | 0.5517 | 0.5557 | 0.5596 | 0.5636 | 0.5675 | 0.5714 | 0.5753 |
| 0.2 | 0.5793 | 0.5832 | 0.5871 | 0.5910 | 0.5948 | 0.5987 | 0.6026 | 0.6064 | 0.6103 | 0.6141 |
| 0.3 | 0.6179 | 0.6217 | 0.6255 | 0.6293 | 0.6331 | 0.6368 | 0.6406 | 0.6443 | 0.6480 | 0.6517 |
| 0.4 | 0.6554 | 0.6591 | 0.6628 | 0.6664 | 0.6700 | 0.6736 | 0.6772 | 0.6808 | 0.6844 | 0.6879 |
| 0.5 | 0.6915 | 0.6950 | 0.6985 | 0.7019 | 0.7054 | 0.7088 | 0.7123 | 0.7157 | 0.7190 | 0.7224 |
| 0.6 | 0.7257 | 0.7291 | 0.7324 | 0.7357 | 0.7389 | 0.7422 | 0.7454 | 0.7486 | 0.7517 | 0.7549 |
| 0.7 | 0.7580 | 0.7611 | 0.7642 | 0.7673 | 0.7704 | 0.7734 | 0.7764 | 0.7794 | 0.7823 | 0.7852 |
| 0.8 | 0.7881 | 0.7910 | 0.7939 | 0.7967 | 0.7995 | 0.8023 | 0.8051 | 0.8078 | 0.8106 | 0.8133 |
| 0.9 | 0.8159 | 0.8186 | 0.8212 | 0.8238 | 0.8264 | 0.8289 | 0.8315 | 0.8340 | 0.8365 | 0.8389 |
| 1.0 | 0.8413 | 0.8438 | 0.8461 | 0.8485 | 0.8508 | 0.8531 | 0.8554 | 0.8577 | 0.8599 | 0.8621 |
| 1.1 | 0.8643 | 0.8665 | 0.8686 | 0.8708 | 0.8729 | 0.8749 | 0.8770 | 0.8790 | 0.8810 | 0.8830 |
| 1.2 | 0.8849 | 0.8869 | 0.8888 | 0.8907 | 0.8925 | 0.8944 | 0.8962 | 0.8980 | 0.8997 | 0.9015 |
| 1.3 | 0.9032 | 0.9049 | 0.9066 | 0.9082 | 0.9099 | 0.9115 | 0.9131 | 0.9147 | 0.9162 | 0.9177 |
| 1.4 | 0.9192 | 0.9207 | 0.9222 | 0.9236 | 0.9251 | 0.9265 | 0.9279 | 0.9292 | 0.9306 | 0.9319 |
| 1.5 | 0.9332 | 0.9345 | 0.9357 | 0.9370 | 0.9382 | 0.9394 | 0.9406 | 0.9418 | 0.9429 | 0.9441 |
| 1.6 | 0.9452 | 0.9463 | 0.9474 | 0.9484 | 0.9495 | 0.9505 | 0.9515 | 0.9525 | 0.9535 | 0.9545 |
| 1.7 | 0.9554 | 0.9564 | 0.9573 | 0.9582 | 0.9591 | 0.9599 | 0.9608 | 0.9616 | 0.9625 | 0.9633 |
| 1.8 | 0.9641 | 0.9649 | 0.9656 | 0.9664 | 0.9671 | 0.9678 | 0.9686 | 0.9693 | 0.9699 | 0.9706 |
| 1.9 | 0.9713 | 0.9719 | 0.9726 | 0.9732 | 0.9738 | 0.9744 | 0.9750 | 0.9756 | 0.9761 | 0.9767 |
| 2.0 | 0.9772 | 0.9778 | 0.9783 | 0.9788 | 0.9793 | 0.9798 | 0.9803 | 0.9808 | 0.9812 | 0.9817 |
| 2.1 | 0.9821 | 0.9826 | 0.9830 | 0.9834 | 0.9838 | 0.9842 | 0.9846 | 0.9850 | 0.9854 | 0.9857 |
| 2.2 | 0.9861 | 0.9864 | 0.9868 | 0.9871 | 0.9875 | 0.9878 | 0.9881 | 0.9884 | 0.9887 | 0.9890 |
| 2.3 | 0.9893 | 0.9896 | 0.9898 | 0.9901 | 0.9904 | 0.9906 | 0.9909 | 0.9911 | 0.9913 | 0.9916 |
| 2.4 | 0.9918 | 0.9920 | 0.9922 | 0.9925 | 0.9927 | 0.9929 | 0.9931 | 0.9932 | 0.9934 | 0.9936 |
| 2.5 | 0.9938 | 0.9940 | 0.9941 | 0.9943 | 0.9945 | 0.9946 | 0.9948 | 0.9949 | 0.9951 | 0.9952 |
| 2.6 | 0.9953 | 0.9955 | 0.9956 | 0.9957 | 0.9959 | 0.9960 | 0.9961 | 0.9962 | 0.9963 | 0.9964 |
| 2.7 | 0.9965 | 0.9966 | 0.9967 | 0.9968 | 0.9969 | 0.9970 | 0.9971 | 0.9972 | 0.9973 | 0.9974 |
| 2.8 | 0.9974 | 0.9975 | 0.9976 | 0.9977 | 0.9977 | 0.9978 | 0.9979 | 0.9979 | 0.9980 | 0.9981 |
| 2.9 | 0.9981 | 0.9982 | 0.9982 | 0.9983 | 0.9984 | 0.9984 | 0.9985 | 0.9985 | 0.9986 | 0.9986 |
| 3.0 | 0.9987 | 0.9987 | 0.9987 | 0.9988 | 0.9988 | 0.9989 | 0.9989 | 0.9989 | 0.9990 | 0.9990 |
| 3.1 | 0.9990 | 0.9991 | 0.9991 | 0.9991 | 0.9992 | 0.9992 | 0.9992 | 0.9992 | 0.9993 | 0.9993 |
| 3.2 | 0.9993 | 0.9993 | 0.9994 | 0.9994 | 0.9994 | 0.9994 | 0.9994 | 0.9995 | 0.9995 | 0.9995 |
| 3.3 | 0.9995 | 0.9995 | 0.9995 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9997 |
| 3.4 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9998 |

