This is a take home midterm. You can use your notes, my online notes on canvas and the textbooks book. You are supposed to work on your own text without external help. I'll be available to answer question in person or via email. Please, write clearly and legibly and take a readable scan before uploading.

To solve the Exam problems, I have not collaborated with anyone nor sought external help and the material presented is the result of my own work.

Name (print): $\qquad$

| Question: | 1 | 2 | 3 | 4 | 5 | Total |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Points: | 10 | 20 | 30 | 20 | 20 | 100 |
| Score: |  |  |  |  |  |  |

Question 1....................................................................................... 10 point
Let $X$ be a continuous r.v. such that

$$
\mathbb{P}(X>0 \& X<M)=1
$$

for some $M>0$. Show that

$$
\mathbb{E}(X)=\int_{0}^{M}\left(1-F_{X}(x)\right) d x
$$

where $F_{X}$ is the c.d.f. of $X$.

## Solution:

Since $\left(1-F_{X}(x)\right)^{\prime}=-f_{X}(x)$ we have

$$
\mathbb{E}(X)=\int_{0}^{M} x f_{X}(x) d x=\left.x\left(1-F_{X}(x)\right)\right|_{x=0} ^{M}+\int_{0}^{M}\left(1-F_{X}(x)\right) d x
$$

Since $F(M)=\mathbb{P}(X<M)=1$ we get the thesis.
Alternatively observe that

$$
\left.1-F_{X}(x)\right)=\int_{x}^{\infty} f_{X}(y) d y=\int_{x}^{M} f_{X}(y) d y d x
$$

so that

$$
\begin{aligned}
\int_{0}^{M}\left(1-F_{X}(x)\right) d x= & \int_{0}^{M} \int_{x}^{M} f_{X}(y) d y d x=\int_{0}^{M} \int_{0}^{y} f_{X}(y) d x d y= \\
& =\int_{0}^{M} y f_{X}(y) d y=\mathbb{E}(X)
\end{aligned}
$$


Let $X_{t}$ describe the price of a stock at time $t$, where time is measured in days and $t=0,1,2, \ldots$ Assume that $X_{t}$ satisfies:

$$
X_{t+1}=\Delta_{t} X_{t}
$$

where $\Delta_{t}, t=0,1,2, \ldots$, form a family of i.i.d. random variables with p.d.f.:

$$
f_{\Delta}(\delta)=\frac{1}{\delta \sqrt{2 \pi \sigma^{2}}} \exp \left(-\frac{(\log \delta-\mu)^{2}}{2 \sigma^{2}}\right) \quad \delta \geq 0
$$

(a) (10 points) Assume that $X_{0}=1$ with probability 1. Find the p.d.f. of $X_{t}$ for $t>0$. (Hint: you can express $\Delta$ is term of a normal r.v..)

Solution: Call $Z_{t}=\log \Delta_{t}$. From the formula of change of variable we get that the p.d.f. of $Z$ is

$$
f_{Z}(z)=\frac{1}{\sqrt{2 \pi \sigma^{2}}} \exp \left(-\frac{(z-\mu)^{2}}{2 \sigma^{2}}\right)
$$

that is $Z$ is normal with $E(Z)=\mu$ and $\operatorname{var}(Z)=\sigma^{2}$. Thus we can write

$$
X_{t+1}=e^{Z_{t}} X_{t}
$$

from which we get $X_{t}=e^{T_{t}}$ where $T_{t}=\sum_{s=0} t-1 Z_{s}$ is a normal r.v. with $\mathbb{E}\left(T_{t}\right)=t \mu$ and $\operatorname{var}\left(T_{t}\right)=t \sigma^{2}$. Changing variables back we get:

$$
f_{X_{t}}(x)=\frac{1}{x \sqrt{2 \pi t \sigma^{2}}} \exp \left(-\frac{(\log x-t \mu)^{2}}{2 t \sigma^{2}}\right) \quad x \geq 0
$$

(b) (10 points) Assume now that $\mu=0.1$ and $\sigma^{2}=0.2$. Find $\bar{x}$ such that

$$
\mathbb{P}\left(X_{10}>\bar{x}\right)=0.75
$$

You can use a calculator like the one here.

## Solution:

We have

$$
\begin{aligned}
\mathbb{P}\left(X_{10}>\bar{x}\right)= & \mathbb{P}\left(T_{10}>\log (\bar{x})\right)=\mathbb{P}\left(\frac{T_{10}-10 \mu}{\sqrt{10} \sigma}>\frac{\log (\bar{x})-10 \mu}{\sqrt{10} \sigma}\right)= \\
& 1-\Phi\left(\frac{\log (\bar{x})-10 \mu}{\sqrt{10} \sigma}\right)=0.75
\end{aligned}
$$

Since $\Phi^{-1}(0.25)=-0.674$ we get

$$
\bar{x}=\exp (-1.414 \cdot 0.674+1)=1.048
$$

## Question 3

30 point
Let $X, Y$ have a joint distribution that is uniform on the unit circle in $\mathbb{R}^{2}$, that is

$$
f_{X, Y}(x, y)= \begin{cases}\frac{1}{\pi} & \text { if } x^{2}+y^{2} \leq 1 \\ 0 & \text { otherwise }\end{cases}
$$

(a) (10 points) Let $R=X^{2}+Y^{2}$ and $\Theta$ be the r.v. defined by

$$
\cos (\Theta)=\frac{X}{\sqrt{X^{2}+Y^{2}}} \quad \sin (\Theta)=\frac{Y}{\sqrt{X^{2}+Y^{2}}}
$$

with $0 \leq \Theta<2 \pi$. Show that $R$ and $\Theta$ are in dependent and find the p.d.f. $f_{R}$ of $R$ and $f_{\Theta}$ of $\Theta$.

Solution: By construction we have

$$
X=\sqrt{R} \cos (\Theta) \quad Y=\sqrt{R} \sin (\Theta)
$$

so that the Jacobian is

$$
J=\left(\begin{array}{cc}
-\frac{\cos (\theta)}{2 \sqrt{r}} & -\sqrt{r} \sin (\theta) \\
-\frac{\sin (\theta)}{2 \sqrt{r}} & \sqrt{r} \cos (\theta)
\end{array}\right)
$$

and $|J|=1 / 2$. Thus we get $f_{R, \Theta}=1 / 2 \pi$ for $0 \leq r \leq 1$ and $0 \leq \theta \leq 2 \pi$. Thus $R$ is uniform in $[0,1]$ while $\Theta$ is uniform in $[0,2 \pi]$ and they are independent.
(b) (10 points) Let $T=\sqrt{-2 \log \left(X^{2}+Y^{2}\right)}$. Fint the p.d.f. $f_{T}$ of $T$.

Solution: We have

$$
\mathbb{P}(T \geq t)=\mathbb{P}\left(X^{2}+Y^{2} \leq e^{-t^{2} / 2}\right)=e^{-t^{2} / 2}
$$

so that

$$
f_{T}(t)=t e^{-t^{2} / 2}
$$

(c) (10 points) Consider the r.v.

$$
\begin{aligned}
U & =\frac{X}{\sqrt{X^{2}+Y^{2}}} \sqrt{-2 \log \left(X^{2}+Y^{2}\right)} \\
V & =\frac{Y}{\sqrt{X^{2}+Y^{2}}} \sqrt{-2 \log \left(X^{2}+Y^{2}\right)} .
\end{aligned}
$$

Show that $U, V$ are i.i.d normal standard r.v.
Solution: We have

$$
\begin{aligned}
U & =\cos (\Theta) T \\
V & =\sin (\Theta) T
\end{aligned}
$$

or

$$
\begin{aligned}
T & =\sqrt{U^{2}+V^{2}} \\
\Theta & =\arctan \left(\frac{U}{V}\right)
\end{aligned}
$$

and we get $|J|=1 / \sqrt{U^{2}+V^{2}}$. Using the change of variable formula we get

$$
f_{U, V}(u, v)=\frac{1}{2 \pi} e^{-\left(u^{2}+v^{2}\right) / 2}=\frac{1}{\sqrt{2 \pi}} e^{-u^{2} / 2} \frac{1}{\sqrt{2 \pi}} e^{-v^{2} / 2}
$$

Question 4 20 point
Let $X_{i}, i=1,2,3, \ldots$, be a family of i.i.d random variables with exponential distribution with parameter 1 , that is

$$
f_{X_{i}}(x)=e^{-x}
$$

For $n \geq 1$ defines $T_{n}=\sum_{i=1}^{n} X_{i}$.
You can think that $X_{i}$ is the time between the $i$-th and the $i+1$-th arrival of a request for service at a certain computer server. In this case, $T_{n}$ is the time of arrival of the $n$-th request for service.
(a) (10 points) Show that the p.d.f. of $T_{n}$ is

$$
f_{T_{n}}(t)=\frac{t^{n-1}}{(n-1)!} e^{-t}
$$

(Hint: Use change of variables to find the joint p.d.f. of the $T_{i}, i=1,2, \ldots, n$, and then compute the marginal on $T_{n}$. Alternatively you can use induction.)

Solution: The joint p.d.f of $X_{1}, X_{2}, \ldots X_{n}$ is $f_{X, n}\left(x_{1}, \ldots, x_{n}\right)=\exp \left(-\sum_{i=1}^{n} x_{i}\right)$ for $x_{i}>0$ and 0 otherwise. Using the formula of change of variables we get that the joint p.d.f. of $T_{1}, T_{2}, \ldots, T_{n}$ is

$$
f_{T_{n}}\left(t_{1}, \ldots, t_{n}\right)=\exp \left(-t_{n}\right)
$$

if $0<t_{1}<t_{2}<\cdots<t_{n}$ and 0 otherwise. Thus we get

$$
f_{T_{n}}\left(t_{n}\right)=\int_{0<t_{1}<\cdots<t_{n}} d t_{1} \cdots d t_{n-1} e^{-t_{n}}=t_{n}^{n-1} e^{-t_{n}} \int_{0<s_{1}<\cdots<s_{n-1}<1} d s_{1} \cdots d s_{n-1}
$$

where we called $s_{i}=t_{i} / t_{n}$. Since

$$
\int_{0<s_{1}<\cdots<s_{n-1}<1} d s_{1} \cdots d s_{n-1}=\frac{1}{(n-1)!}
$$

we get

$$
f_{T_{n}}\left(t_{n}\right)=\frac{t_{n}^{n-1}}{(n-1)!} e^{-t_{n}}
$$

Observe that $T_{n+1}=T_{n}+X_{n+1}$ and $T_{n}$ is independent from $X_{n+1}$ so that

$$
f_{T_{n+1}}(t)=\int_{0}^{\infty} f_{X}(t-s) f_{T_{n}}(s) d s
$$

thus assuming that $f_{T_{n}}$ satisfy the thesis we get

$$
f_{T_{n+1}}(t)=\int_{0}^{t} e^{-(t-s)} \frac{s^{n-1}}{(n-1)!} e^{-s} d s=\frac{e^{-t}}{(n-1)!} \int_{0}^{t} s^{n-1} d s=\frac{t^{n}}{n!} e^{-t}
$$

(b) (10 points) Let $N_{t}$ be the number of arrivals before time $t$, that is

$$
N_{t}=\max \left\{n \mid T_{n}<t\right\} .
$$

Show that $N_{t}$ is a Poisson r.v. with expected value $t$. (Hint: express $\mathbb{P}\left(N_{t}<n\right)$ in term of $T_{n}$.)

Solution: Observe that $\mathbb{P}\left(N_{t}<n\right)=\mathbb{P}\left(T_{n} \geq t\right)$ so that

$$
\begin{aligned}
\mathbb{P}\left(N_{t} \leq n\right)= & \int_{t}^{\infty} \frac{s^{n}}{n!} e^{-s} d s=\int_{0}^{\infty} \frac{(t+s)^{n}}{n!} e^{-(t+s)} d s= \\
& \sum_{k=0}^{n}\binom{n}{k} \frac{t^{k} e^{-t}}{n!} \int_{0}^{\infty} s^{n-k} e^{-s} d s=\sum_{k=0}^{n} \frac{t^{k} e^{-t}}{k!}
\end{aligned}
$$

so that

$$
\mathbb{P}\left(N_{t}=n\right)=\mathbb{P}\left(N_{t} \leq n\right)-\mathbb{P}\left(N_{t} \leq n-1\right)=\frac{t^{n} e^{-t}}{n!}
$$

Question 5 20 point
Two question on continuous/discrete r.v.
(a) (10 points) Let $X$ be continuous r.v. uniformly distributed in $[0,1]$. Consider the r.v.

$$
Z=\max \{X, 0.5\}
$$

Find the c.d.f. of $Z$. Is $Z$ a continuous r.v.? Is it discrete?

Solution: Clearly we have $\mathbb{P}(Z<0.5)=0$ while $\mathbb{P}(Z \leq z)=\mathbb{P}(X \leq z)$ if $z \geq 0.5$ so that we have

$$
F_{Z}(z)= \begin{cases}0 & z<0.5 \\ z & 0.5 \leq z<1 \\ 1 & z \geq 1\end{cases}
$$

$Z$ is not a continuous r.v. since $F_{Z}$ is not continuous and it is not discrete since $F_{Z}$ is not piecewise constant.
(b) (10 points) Let $X$ be continuous r.v. with p.d.f. $f_{X}$ and $Y$ be a discrete r.v. with p.m.f. $p_{Y}$. Moreover $X$ and $Y$ are independent. Find the p.d.f. of $Z=X+Y$. Is $Z$ a continuous r.v.?

Solution: We have

$$
\mathbb{P}(Z \leq z)=\sum_{y} \mathbb{P}(X \leq z-y \& Y=y)=\sum_{y} \mathbb{P}(X \leq z-y) \mathbb{P}(Y=y)
$$

so that

$$
f_{Z}(z)=\sum_{y} f_{X}(z-y) p_{Y}(y)
$$

In general we cannot say whether it is continuous but if $Y$ takes only finitely many values than $Z$ is continuous.

